# REVIEW



# Does residential address-based exposure assessment for outdoor air pollution lead to bias in epidemiological studies?

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# Abstract

**Background** Epidemiological studies of long-term exposure to outdoor air pollution have consistently documented associations with morbidity and mortality. Air pollution exposure in these epidemiological studies is generally assessed at the residential address, because individual time-activity patterns are seldom known in large epidemiological studies. Ignoring time-activity patterns may result in bias in epidemiological studies. The aims of this paper are to assess the agreement between exposure assessed at the residential address and exposures estimated with time-activity integrated and the potential bias in epidemiological studies when exposure is estimated at the residential address.

**Main body** We reviewed exposure studies that have compared residential and time-activity integrated exposures, with a focus on the correlation. We further discuss epidemiological studies that have compared health effect estimates between the residential and time-activity integrated exposure and studies that have indirectly estimated the potential bias in health effect estimates in epidemiological studies related to ignoring time-activity patterns.

A large number of studies compared residential and time-activity integrated exposure, especially in Europe and North America, mostly focusing on differences in level. Eleven of these studies reported correlations, showing that the correlation between residential address-based and time-activity integrated long-term air pollution exposure was generally high to very high (R > 0.8). For individual subjects large differences were found between residential and time-activity integrated exposures. Consistent with the high correlation, five of six identified epidemiological studies found nearly identical health effects using residential and time-activity integrated exposure. Six additional studies in Europe and North America showed only small to moderate potential bias (9 to 30% potential underestimation) in estimated exposure response functions. Exposure contrasts were smaller for time-activity integrated exposures in nearly all studies. The difference in exposure was not equally distributed across the population including between different socio-economic groups.

**Conclusions** Overall, the bias in epidemiological studies related to assessing long-term exposure at the residential address only is likely small in populations comparable to those evaluated in the comparison studies. Further improvements in exposure assessment especially for large populations remain useful.

Keywords Mobility, Static exposure, Dynamic exposure, Epidemiology, Personal exposure, Air pollution

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### Background

Air pollution has been associated with a range of adverse health effects [59]. Outdoor air pollution is the most important environmental exposure in the Global Burden of Disease assessments [8]. The evidence for these burden of disease assessments derives primarily from epidemiological studies, especially on long-term air pollution exposure that is particularly important in terms of health. A Panel appointed by the Health Effects Institute (HEI) has recently reviewed epidemiological studies of traffic-related air pollution, addressing exposure assessment issues in detail [4, 18]. The Panel primarily assessed how well specific methods assessed the outdoor concentration and how well outdoor exposure was assigned to the residential address, e.g. by assessing the spatial resolution of the exposure surfaces and addresses. Virtually all studies assigned outdoor concentrations to the residential address only [18], a few studies in children incorporated exposures at the school address. Even fewer studies in adults have incorporated work address in the exposure assessment.

Environmental health researchers have understood for decades that the true personal exposure to air pollution is experienced in multiple so-called micro-environments [15, 41]. Personal exposure can be assessed by direct personal exposure monitoring or indirectly by assessing concentrations in key micro-environments and obtaining time-activity data [15]. There is a large exposure science literature on both approaches. However, in large epidemiological studies neither direct nor indirect personal exposure have often been assessed. The main reason for this is that personal exposure monitoring is too costly to perform in a large number of subjects and that assessing long-term exposure requires a fair number of repeated samples per subject. Indirect exposure assessment is also not applied often, because most epidemiological studies do not have information on where people spend time other than the home location. Post-hoc time-activity diaries are also prohibitive in large cohort studies and not feasible in retrospective cohort studies. The lack of timeactivity data beyond the residential address is related to the fact that very few cohort studies, which have been used in air pollution epidemiology, have been primarily designed to investigate health effects of air pollution. In the recent large studies based on administrative databases, information on residential address is also typically the only location data available [13, 55]. The few studies that were designed to investigate outdoor air pollution and did obtain more detailed data on work or school address, include the SAPALDIA study in Switzerland and the PIAMA birth cohort study in the Netherlands [32, 61].

Epidemiological studies on air pollution and other environmental exposures have been criticized for not taking time-activity patterns into account [24, 42, 56]. Residence-based exposure assessment is generally criticized because of the fact that air pollution differs in space and people spend a sizable amount of time outside their home. This is occasionally documented with anecdotal figures of individual subject's exposure data along specific tracks, showing large variability depending on location and time [58]. We note that this alone is insufficient to make strong statements about potential bias in epidemiological studies as the documented exposure contrasts are not the contrasts used in epidemiological studies: what is used is the contrast in long-term average (e.g., annual average) concentrations across individuals. Furthermore, major air pollutants including fine particles have been shown to infiltrate efficiently into homes [20]. While it is understandable that epidemiological studies have relied on the residential address, the questions remain as to how poorly exposure is assessed and how much bias is potentially introduced in health effect estimates by focusing on the residential address only.

The objective of this paper is to review studies that have evaluated how air pollution exposures assessed at the residence only compare with exposures that integrate time-activity patterns. Our second objective is to review studies that have evaluated the potential bias of using residential exposure assessment in epidemiological studies of long-term air pollution exposure.

### Main text

# Methods

### **Review methods**

We built on a previous review discussing methods for long-term air pollution exposure assessment in which the issue of residential exposure assessment was addressed [19]. We added more recent studies by searching in the database Pubmed with the search terms "mobility", "static exposure", "dynamic exposure", "personal exposure", "residential exposure" AND "air pollution" OR "air pollutants". We limited the search to studies in humans and in the English language. The search was conducted February 15, 2024. In addition, we reviewed the reference lists of identified papers. We treat the following exposurerelated terms used by researchers in the identified papers as synonymous: a) dynamic, mobility enhanced, integrated, time-activity integrated and personal exposure, and b) static, home-based and residential exposure.

# Methods to assess time-activity patterns and dynamic exposures

We can compare residential exposure with directly measured personal exposure or indirectly assessed personal

exposure based on validation studies. The literature using directly measured personal exposure as a validation metric is small and difficult to interpret [35, 38]. In these studies, statistically significant but weak associations between long-term outdoor and personal exposure were found. The main problem for the interpretation is the difficulty in obtaining a sufficiently large number of measurements per person for a large group of subjects to credibly assess long-term personal exposure. A second reason is that measured personal exposure of e.g., PM<sub>2.5</sub> and NO<sub>2</sub> includes both outdoor and indoor sources, which are difficult to separate. This is in sharp contrast to personal exposure validation studies of short-term air pollution, which have convincingly documented that the temporal variation at a single outdoor site correlates well with temporal variations in personal exposure [5]. The spatial comparison is much more complicated, see the IARC monograph on outdoor air pollution for a discussion [20]. In the remainder of this review, we focus on studies comparing residential exposure with indirectly assessed personal exposure.

Obtaining information on time-activity patterns of individuals to calculate personal exposure as a weighted sum of time spent in different micro-environments, and the concentration in that micro-environment, is also challenging. Key micro-environments for air pollution exposure are the residence, work or school location and commuting route [10]. We identified three approaches to assess time-activity patterns that have been applied in the framework of exposure assessment and epidemiological studies. First, agent-based modelling has been applied to simulate individual time-activity, typically based upon existing survey data [29, 37]. In most applications, the actual individual work or school address is not known. This approach can be applied to large populations. Second, tracking studies have been performed in which volunteers carried a GPS tracker or smartphone App to empirically determine time-activity patterns [30, 31, 60]. Tracking studies are typically studies in small population groups (several hundred at most) and cover a limited amount of time, often one to two weeks because of the demand on the participants' time. Third, a few epidemiological studies, designed to assess environmental exposures such as air pollution and noise, have collected information on the school or work address, including a large Canadian cohort using census data [7] and a study in Southern-Californian children [33]. While these studies did not formally obtain full time-activity patterns, notably not capturing mobility between these locations and leisure time activities, we included these studies as school and work cover major micro-environments for children and working-age adults.

Studies have compared residence-based exposure and indirectly assessed personal exposure in different ways. Some studies compared the level and contrast of exposure [29, 31]. One of the first empirical tracking studies compared the contribution of the residential environment to the time-integrated exposure with that of other micro-environments [10]. As the primary focus of our review is application in epidemiological studies, an important comparison metric is the *correlation* between residence-based exposure and indirectly assessed personal exposure. The correlation is a very important metric because in epidemiological studies we compare the health status of individuals with high and low exposures, typically on a continuous scale. If the correlation is low, we may anticipate different associations with health depending on whether exposure is assessed at the residence only or with time-activity integrated. In the case of classical exposure measurement error, we typically expect a downward bias of effect estimates in epidemiological studies [49]. We included studies if they reported correlations or directly assessed the possible bias in epidemiological studies. We additionally evaluated the difference in contrasts and absolute levels between residence-based and indirect personal exposure. In the case of a high correlation between the two exposure estimates, higher health effect estimates may be obtained when using an exposure metric with a smaller contrast. Finally, we evaluated the epidemiological studies that directly compared health effect estimates using residential and time-activity integrated exposures.

### Time-activity surveys

Time-activity surveys have shown that people spend about 60–70% of their time in their own home environment, with differences between population groups and individuals [20, 26]. In two large North American surveys (Canada and USA), people (including adults and children) spent 65–66% in the home, 10% at work/school and 6% outdoors [26]. Children spent about 6–8% more time at home in both cohorts [26]. In a large German survey of children, children spent 15.5 h per day at home, 4.75 in other indoor locations and 3.6 h outdoors [9]. In a large survey of 100,000 people in Los Angeles county, USA, the population spent 66% of their time in their home (workers 60% and non-workers 72%) [28].

### Results

Figure 1 shows the selection of studies. The search resulted in 2139 abstracts to be screened, of which the full text of 107 were assessed for eligibility. In total, 29 studies were included in the review for quantitative comparisons. Exclusions were related to No comparison of residence-based and dynamic exposure (n=45), Papers

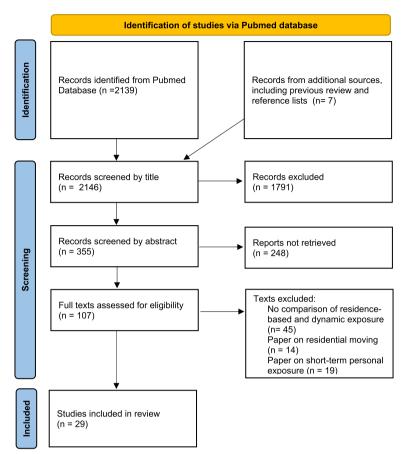


Fig. 1 Flowchart of selection of studies for quantitative comparison of residence-based and dynamic air pollution exposure

addressing residential moving in typically birth cohort studies (n = 14) and papers on short-term personal exposure (n = 19). Of the 45 excluded studies, some were used in the supporting text.

# Comparison between residential and time-activity integrated exposure level and contrast

Table 1 shows the identified studies that analyzed the relationship between residential address-based exposure and time-activity integrated exposure assessment. Studies were primarily conducted in North America and Europe, with three studies from China.

*Exposure levels* were generally only modestly different between residential and dynamic exposure assessment methods on a population average level, with residential exposure being higher in some studies and lower in other studies (Table 1). Exposure was on average *higher* in Canada when the work address was incorporated in the exposure assessment compared to the residential address alone, as work addresses were within locations that on average had higher PM<sub>2.5</sub> concentrations [7]. A study in the New York City region, USA, showed significantly different population weighted PM<sub>2.5</sub> exposures in 71 districts between active and home scenarios, although differences were small in absolute terms (-0.26 to 0.73 µg/  $m^{3}$ ), [39]. In studies in Boston, USA and the Netherlands, mean levels were similar for static and dynamic exposures [36, 40]. In a study in Montreal, Canada, almost 90% of the population had a higher integrated NO<sub>2</sub> exposure compared to the residential exposure [52]. In another study in Montreal, NO<sub>2</sub> and UFP exposures were lower for dynamic exposures of UFP and NO<sub>2</sub> and nearly identical for PM<sub>2.5</sub> [16]. Substantially larger mobilityintegrated exposures compared to residential exposures were found for UFP and BC in Toronto, Canada related to commuting exposures [53]. Residential exposures of 106 participants in a study in Beijing were higher compared to mobility-based exposures, both on low and high pollution days [30]. All subjects lived in a highly polluted neighborhood and therefore mobility resulted in lower mobility-integrated exposures [30]. A study in Hong-Kong found lower dynamic exposures for PM<sub>2.5</sub> and BC and higher dynamic exposures for  $NO_2$  [58]. A study in the UK found a small increase in population weighted NO<sub>2</sub> and PM<sub>2.5</sub> exposures when including exposure at the work address (2 and 0.3% respectively) [46].

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Table 1
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Population	Study area	Dynamic EA method	Pollutants	Correlation	Mean level and contrast (μg/m <sup>3</sup> )	Reference
<b>Studies reporting correlation</b> Participants of survey (N = 680)	Cantons of Basel-Stadt and Baselland, Switzerland	Actual individual data on home, work and com- mute trip	NO <sub>2</sub> from dispersion model	Correlation between RES and PERS (HWC) 0.97. RES exposure correlated with work ( $R = 0.48$ ) and commute ( $R = 0.69$ )	Mean (SD): RES 25.6 (6.3) HW 26.1 (5.7) PERS (HWC) 26.3 (5.6)	Ragettli, 2014 [45]
Most school children (950,000) 6–18 yr and a rep- resentative sample of 24 – 65 yr adults (N~350,000)	Israel central coastal plain	Known individual home, work and school address to estimate PERS exposure. Shortest route modelling between home and school / work	NO <sub>2</sub> from dispersion model	Correlation reported as R <sup>2</sup> between RES and PERS 0.64 for adults and 0.89 for school children. Correlations lower for commuters	NR, but average difference between home and school / work exposure close to zero	[20]
Pregnant women (N=486)	Paris, France	Questionnaire on time at work, commuting, trans- port mode	NO <sub>2</sub> assessed at the census block scale	Correlation between RES and PERS 0.94	The difference between PERS and RES was on average 1.46 (~ 3%). The difference was higher in deprived neighborhoods	[]]
Adults (N= 9,886)	Shenzhen, China	Cell-phone location data	PM <sub>2.5</sub> , NO <sub>2</sub> , EC at 3*3 km for a single weekday in Octo- ber	High correlations between RES and activity- based: R <sup>2</sup> 0.76, 0.74 and 0.69 for PM <sub>2.5</sub> , NO <sub>2</sub> and EC. Large differences for individuals	Mean (SD) for NO <sub>2</sub> activity- based and RES 23.2 (2.74) and 23.0 (3.44); PM <sub>2.5</sub> 72.9 (4.44) and 7.2.5 (5.53); EC 10.7 (0.88) and 10.6 (1.23)	[62]
Synthetic population	Utrecht city, the Netherlands	ABM, incorporating work and commuting	NO <sub>2</sub> from an hourly LUR model based on routine monitoring sites	Very high correlations: $R^2 = 0.93$ between RES and PERS and $R^2 = 0.98$ between HOMEMAKER and PERS	Median (IQR) 20.6 (2.9) for homemaker and 21.4 (1.9) for commuters	[29]
Children 5 yr of Whistler cohort (N=733)	Neighborhood in Utrecht, the Netherlands	ABM, incorporating school, commuting to school and playing outdoors. Time fractions and distance to locations from surveys	NO <sub>2</sub> , PM <sub>2.5</sub> , PM <sub>2.5</sub> absorbance from LUR	RES and PERS highly correlated: $R = 0.84$ for NO <sub>2</sub> ; larger for BC and PM <sub>2.5</sub>	Mean (SD) for NO <sub>2</sub> 28.9 (2.7) for RES and 29.4 (2.1) for PERS; for $PM_{2.5}$ 16.7 (0.3) and 16.7 (0.2); for BC 1.3 (0.1) and 1.3 (0.1)	[37]
Adults (N= 310,989)	Shenzhen, China	Cell-phone location data	PM <sub>2.5</sub> , NO <sub>2</sub> , CO, NO2, SO2, at 3*3 km by CMAQ and inverse distance weigh- ing	High correlations between RES and activity- based: $R^2$ for PM <sub>25</sub> and NO <sub>2</sub> for CMAQ model. Higher for interpolation	Mean NO <sub>2</sub> 16.1 and 16.1 ppb for RES and PERS for the low- est mobility and 15.6 and 15.1 ppb for the highest mobility group	Yu, 2018 [62]
Children 8 yr of PIAMA cohort (N=668)	Nationwide in the Nether- lands	ABM, incorporating school, commuting to school and playing outdoors. Time fractions and distance to locations from surveys	NO <sub>2</sub> , PM <sub>25</sub> , PM <sub>25</sub> absorbance from LUR	Very high correlation (R > 0.9) between RES and PERS for PM2.5, NO2 and BC. Same for estimated and actual school	Mean (SD) for NO <sub>2</sub> 22.9 (6.6) for RES and 23.7 (6.5) for PRS; for $PM_{25}$ 16.3 (0.6) and 16.3 (0.6) and 16.3 (0.6); for BC 1.2 (0.2) and 1.3 (0.2)	[36]

Population	Study area	Dynamic EA method	Pollutants	Correlation	Mean level and contrast (µg/m³)	Reference
Tracking panel adults (N= 269)	The Netherlands	Activity-space based assess- ment integrating actual GPS readings	PM <sub>2.5</sub> (and noise, green and blue space) from LUR	Very high correlation between RES and activity- space based exposure (R 0.83 – 0.94 for different activ- ity spaces)	Contrast smaller for RES defined as 1000 m buffer, IQR= 1.92 versus 1.5 -1 .6 for different activity-space based	[60]
Full population includ- ing children and adults	London, UK	Time survey data, account- ing for indoor-outdoor differences and commuting. 95–98% indoors	PM <sub>25</sub> and NO <sub>2</sub>	Correlation between RES and dynamic dependent on commuting: R $0.77$ and $0.79$ for active travelers and $0.39$ and $0.56$ for inactive travelers for PM <sub>2.5</sub> and NO <sub>2</sub>	Average concentration 37% an $63\%$ lower for dynamic exposure for $PM_{2,5}$ and $NO_2$ respectively (related to indoor / outdoor ratio of 0.56 and 0.31)	[54]
Tracking panel adults The including depression analysis (N = 393)	The Netherlands	Activity-space based assess- ment integrating actual GPS readings for 7 days	PM <sub>2.5</sub> from LUR	R between RES and PERS 0.80 – 0.84 for different buffers	Mean (SD) 16.6 (0.7) and 17.0 (1.2) for RES and PERS, based on 50 m buffer around home	[47]
Adults of the CAFEH study, living near (≤ 500 m) or farther away (≥ 1000 m) from a major highway (N= 140)	Boston, USA	Time-activity data (15 min resolution) from question- naire of two days, hourly modelling	UFP from a spatio-temporal LUR model based on mobile monitoring	Ж Х	Mean (SD) 25,000 (4000) for RES and 23,000 (3000) for PERS	[25]
Total population of NYC (N=8.5 million)	NYC region, USA	Mobile device usage records defining the active population in a district by hour of the day	PM <sub>2.5</sub> from interpolation of 155 monitors across NYC assigned to district	NR	Small but significant differ- ences between population- weighted average of RES versus Active of $PM_{2,5}$ in the 71 districts (-0.26 to 0.73 ug/m <sup>3</sup> )	[39]
Participants of a travel survey (N= 5,945)	Montreal, Canada	Origin destination survey of 15,572 trips from 5,945 individuals, 24-h average for 4 weeks across 4 seasons	NO <sub>2</sub> from a dispersion model	N	Mean of RES vs PERS 6.4 vs 6.3, 6.9 vs 6.4 and 6.4 vs 6.2 for drivers, transit users and active transport com- muters. SD differs by mode, but reflects 24-h	[52]
Users of the largest mobile phone network (N~5 mil- lion)	Belgium	Mobile phone location data for two days in October 2015 (weekday and weekend)	NO <sub>2</sub> from an integrated model	NR (scatterplots of RES vs PERS show high correlation)	Mean (SD) NO <sub>2</sub> 29.7 (12.0) and 31.0 (11.3) for RES and PERS exposure for week- day and 27.5 (8.6) and 27.6 (8.0) for weekend	[12]

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Table 1 (continued)						
Population	Study area	Dynamic EA method	Pollutants	Correlation	Mean level and contrast (µg/m <sup>3</sup> )	Reference
Participants of survey (N=89,385)	Hong Kong, China	Agent-based modelling	$PM_{2,5}$ , BC and NO <sub>2</sub> from LUR models	R	Mean (SD) for RES vs PERS for PM <sub>2,5</sub> 33.8 (3.9) and 32.0 (3.4) BC 9.6 (3.4) and 9.4 (3.5) NO <sub>2</sub> 90.9 (18.1) and 92.9 (15.1)	[58]
Participants of the MTL Trajet study (N = 5,452)	Montreal, Canada	Mobile phone GPS data for an average 16 days	PM <sub>25</sub> 1*1 km from satellite data; UFP and NO <sub>2</sub> at 100* 100 m from LUR based on mobile monitoring		Mean for RES vs PERS for UFP 26,500 and 24,780; for $PM_{25}$ 8.69 and 8.74; for NO <sub>2</sub> 23.35 and 22.43. Range smaller for PERS e.g., for NO <sub>2</sub> 41 vs 47 ppb	[16]
Representative sample of working Israeli adults (N~168,000)	Israel central coastal plain	Shortest route modelling between known home and work address to esti- mate PERS exposure	NO <sub>2</sub> from dispersion model	X	Differences between PERS and RES exposure small (<5 pph.hr <sup>-1</sup> ) for commute (short time), leisure time (close to home); larger for working hours (long period). Integrated exposure not reported	[51]
Mobile phone users of two companies (N = 407,435)	Greater Boston area, USA	Anonymous mobile phone data used to infer home and work location	PM <sub>2.5</sub> at 400*400 m hourly resolution from LUR and sat- ellite	NR	Mean and SD were 8.62 (0.76) for RES and 8.6 (0.67) for PERS (combined HW)	[40]
Participants tracking study in high pollution area (N=106)	Beijing, China	GPS sensors and air pollu- tion measurement devices for 48 h (weekday and week- end day) + TAD questionnaire	PM2.5, using kriging on 35 stations in Beijing for RES; PERS from personal monitors	NК	Mean (5D) for RES vs PERS 29.27 (12.7) and 23.07 (11.7) (low pollution days); 83.82 (20.01) and 73.23 (18.23) (high pollution days)	[30]
Participants from a travel survey data (N= 341,274)	Toronto, Canada	Stochastic multi-path assignment model (EMME4) informed by travel survey data from 341,274 individu- als	UFP and BC LUR models using fixed and mobile data	Я	UFP: median PERS 11.6% higher than RES (median 23,784 particles vs 21,316 particles/cm <sup>3</sup> ). BC: median PERS 63.2% higher than RES (median 1321 vs 809 ng/m <sup>3</sup> )	[53]
Canchec 2001 census	Nationwide in Canada	Time weighted average of actual work and home location	PM <sub>3.5</sub> at 1*1 km, MAPLE satellite -based method	ц	Exposure at work higher than residential. Mean and IQR of RES and PERS similar in full cohort. In com- muting cohort, mean (IQR) were 7.3 (3.6) and 7.5 (3.4) for RES and PERS	E

Population	Study area	Dynamic EA method	Pollutants	Correlation	Mean level and contrast (µg/m³)	Reference
Random sample of residents (N=100,789)	Los Angeles County, USA	Agent-based mod- eling on time survey data in Southern-California	PM <sub>2.5</sub> at 500*500 m resolu- tion using Random Forest model based on low-cost sensor monitoring	ж	Mean (SD) 8.1 (1.4) and 8.2 (1.4) for RES and PERS expo- sure for whole population. Difference in mean larger for workers (0.1) versus non- workers (0.01)	Lu, 2021 [28]
Senior citizens ( $n = 113$ )	Barcelona, Spain	GPS tracking	PM <sub>10</sub> , PM <sub>2.5</sub> , NO <sub>2</sub> , noise and green	NR	Small differences in mean concentrations of PM <sub>10</sub> . PM <sub>15</sub> , and NO <sub>2</sub> between RES and PERS (1, 6 and 1% higher for PERS)	[31]
Sample of workers from administrative database (N= 130 million)	United States	Combination of home and work address	PM <sub>2.5</sub>	NR	Mean RES and WORK overall 7.8 and 8. Disparities differed by race and income	[11]

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Almost none of the studies incorporated differences between indoor and outdoor concentrations, related to infiltration into homes, schools or offices. A study in Hong-Kong found that if infiltration indoors was taken into account, dynamic exposures were on average about 20% lower than residential exposures because infiltration of outdoor air pollution was more efficient in residences than in commercial buildings [58]. A study in London found that dynamic exposure accounting for infiltration and commuting was lower than residential outdoor concentrations [54]. The mean dynamic exposure to outdoor sources was estimated to be 37% lower for PM25 and 63% lower for NO<sub>2</sub> than at the residential address. These decreased estimates reflect the effects of reduced exposure indoors (with mean infiltration assumed to be 0.31 for  $NO_2$  and 0.56 for  $PM_{2.5}$ ), the large amount of time spent indoors (95-98%) and the mode and duration of travel in London [54]. A study in schoolchildren in five European cities, found that after taking into account time-activity patterns, exposures at school and in traffic and infiltration, residential outdoor PM2.5 concentrations were on average 26% higher than the integrated exposure [23]. In the models, infiltration factors were 0.66 and 0.82 for residences and schools in the four South-European cities and 0.55 for residences for the Finish city. Children spent 87% and 91% indoors in the South-European cities and Kuopio, Finland respectively [23]. A study in Beijing, reported an annual mean outdoor  $PM_{2.5}$ concentration of 87.6  $\mu$ g/m<sup>3</sup> versus a 47.5  $\mu$ g/m<sup>3</sup> average personal exposure, calculated based on an assumed mean infiltration factor of 0.47 (Shi, 2017). Studies furthermore did not take indoor sources into account, which is reasonable if the interest is in outdoor-generated pollution.

Overall, differences between studies are related to differences in locations where people live and work. Commuting exposures are, in general, especially higher for traffic-related pollutants such as  $NO_2$  and UFP, but most people spend only a small amount of time in traffic. A study in Israel found that the contribution of high  $NO_2$ concentration during commute to the overall, integrated exposure, was small [51].

Variability in dynamic exposure levels was lower compared to residential based exposure in almost all studies (Table 1). In most studies the difference in exposure contrast was small to moderate. The explanation for the smaller dynamic exposure contrast is probably related to the phenomenon that subjects living at the highest exposed locations likely work at less exposed locations and vice versa. Kim and Park (2021) argued that this pattern is common and use the term neighborhood effect averaging, distinguishing upward and downward averaging [22]. Subjects living in the suburbs or towns near the city are more likely to work in higher polluted urban centers. The reverse pattern is less likely, however subjects living in local hotspots such as major roads may work in more urban background conditions. In the Montreal study, plots of the difference between the dynamic and residential exposure show that dynamic exposure was higher for those with low residential exposure levels and lower for those with high residential concentrations [16]. In the NYC study, boxplots comparing residential and integrated exposures by district reveal that the majority of districts have a lower integrated exposure compared to residential exposure, except for one district with a very low residential exposure and the highest integrated exposure [39].

Several studies have investigated determinants of the difference between dynamic and residential exposures. In a study in Paris, France, the difference between dynamic and static exposure was larger for deprived neighborhoods [3]. The locations where people work likely differ by socio-economic group, e.g., because offices and industrial facilities are located in different places. Exposure contrasts between urban and rural areas were different for residential exposure and time-weighted average exposure including the work address [7]. The exposure contrast between major streets and urban background in the city of Utrecht was smaller for dynamic exposure than for residential exposure [29]. In a detailed analysis of residential and mobility-enhanced exposures of UFP, NO<sub>2</sub> and PM<sub>2.5</sub> in Montreal, Canada, various built-environment characteristics (e.g. income, unemployment rate, major road length, walkability) at the neighborhood level and smaller predicted differences for individuals, though with only small differences in the overall average between residential and mobility exposures [16]. In a study in Los Angeles, USA, the overall average static and dynamic PM<sub>2.5</sub> exposure was identical, but for individual subjects larger differences occurred e.g. for those living in a low pollution neighborhood and working in a high pollution area and vice versa, related to socio-economic factors [28]. The difference was larger for those travelling more than 20 miles between home and work, if these two locations differed in exposure [28]. In a large study in US workers, disparities between different population groups by race, income, age and urbanicity were larger when the work address was taken into account [11].

Several studies have found differences in population exposure distributions across neighborhoods or districts depending on using residential or dynamic exposures, related to where people predominantly work and live [39]. In studies in Belgium and the UK, researchers found that NO<sub>2</sub> exposure differed when considering the working-hour and night-time population distribution [2, 14, 46]. In the UK study, population-weighted NO<sub>2</sub>, PM<sub>2.5</sub> and O<sub>3</sub> exposures were on average 2%, 0.3% and -0.3% higher when the workday-population distribution was taken into account [46]. Population exposures differed between daytime and nighttime, partly because many areas are predominantly residential or areas where a large number of people work.

These observations may be very useful for overall population exposure assessment and subsequent risk or health impact assessment. However, in epidemiological studies we need to know the exposure of individuals, and especially the correlation between the residential and more refined exposure estimates incorporating time-activity patterns.

# Correlation between residential and time-activity integrated exposures

Eleven studies reported on the correlation between residential and time-activity integrated exposure. The correlation between residential address-based and dynamic long-term air pollution exposure was generally high to very high (Table 1). This holds both for studies that simulated the time spent in locations other than the home based on agent-based modelling and for studies that had access to the actual individual time-activity patterns. Studies have been performed in adults and children and in different settings (Netherlands, France, Canada, Switzerland, USA, United Kingdom and China). While the overall correlation is high for the full population, large differences may occur for individual subjects as illustrated in the Montreal study [16]. The agreement between residential and mobility-enhanced exposure was less for UFP than for NO<sub>2</sub> and PM<sub>2.5</sub> in the Montreal study (correlation coefficient not reported). In a study in Israel, the correlation was high for working-age adults ( $R^2 = 0.64$ ) and very high for school children ( $R^2 = 0.89$ ) [50]. In the Israel study, the correlation was lower for commuters than for non-commuters:  $R^2 = 0.56$  vs  $R^2 = 0.68$  for the adults and  $R^2 = 0.86$  vs  $R^2 = 0.94$  for the children [50].

In London, correlations differed between commuting mode, despite the small percentage of time spent in travel (1–4%) and the large percentage of time indoors (95–98% across populations) [54]. Across the population, high correlations were found between dynamic exposure including infiltration into homes and offices and residential exposure for non-commuters and those using active travel (R=0.91 and 0.77 for NO<sub>2</sub>), while only moderate correlations were reported for those using inactive travel (R=0.57 for NO<sub>2</sub>, affected by the high concentrations assumed in the London underground) [54].

The generally high correlations between residencebased exposure and time-activity integrated exposure is probably due to the relatively large time fraction spent at home and the correlation between exposure at home and other micro-environments.

One of the first empirical tracking studies compared the contribution of the residential, work and commuting environment to overall exposure [10]. In 36 subjects in Barcelona, Spain the tracking-based study documented that subjects spent about 70% of their time in their residence, and the residence contributed 54% to the personal exposure to NO<sub>2</sub>. In a large study in Southern California, USA, people spent 66% of their time in their own home, which contributed 61% of the overall PM<sub>2.5</sub> exposure [28].

Several studies reported that the exposure at the residence correlated with the exposure at school or work, the two other major micro-environments where people spend time [36, 45]. In the Swiss study including 680 adults and children in the cantons of Basel-Stadt (i.e. "Basel city") and Baselland (i.e. "Basel country" with towns, villages and rural areas), the correlation between residential and work and commuting exposure were 0.48 and 0.69 respectively. In the Dutch study of school-age children, correlations for NO2 between residential and estimated school and travelling exposures were 0.92 and 0.92 (0.90 and 0.88 for  $PM_{25}$ ) [36]. The correlation between residential and actual school exposures was nearly identical (0.89). Therefore, the residential exposure estimate also contains some information on the exposure in other locations, because children tend to go to a school near their home. Adults also on average tend to work relatively close to their home. Given the moderate correlation, the residential exposure clearly does not perfectly predict exposure elsewhere, but it does add information in addition to the time actually spent at home.

In the PIAMA cohort study, information about the actual school address was available. The estimated and actual exposure at school were highly correlated (R=0.93) [36], providing support for the estimation methods. The 24-h integrated exposure incorporating the actual and the integrated exposure using the assumed school was very highly correlated (R=0.995).

# Comparison of health effect estimates in epidemiological studies

Table 2 shows epidemiological studies comparing health effect estimates of residential and time- activity integrated exposures. Consistent with the high correlation between residential and time-activity integrated exposures, health effects estimated from both approaches were nearly identical in five of the six studies. The sixth study used a difficult to interpret exposure metric (exposure in a 1600 m buffer) to represent residential exposure.

In a large Canadian administrative cohort with over ten million person-years of follow-up, significant associations of long-term  $PM_{2.5}$  exposure with natural

Population	Study area	Dynamic EA method	Pollutants	Findings epidemiology	Reference
Southern California Children's Health Study (N=2497 children)	Southern California, USA	Home and school combined, using assumed time at school (21%) and the remainder assigned to home	NO <sub>x</sub> referred to as traffic-related air pollution from a line dispersion model	HR for incident asthma per IQR for home 1.32 (1.08–1.6.1), school 1.20 (0.91–1.58) and combined home and school 1.34 (1.07–1.68)	[33]
Adults of the CAFEH study liv- ing near (≤ 500 m) or farther away (≥ 1000 m) from a major highway (N = 140)	Boston, USA	Time-activity data from questionnaire, hourly modelling	UFP from hourly LUR	Weak associations with residential and dynamic exposure, with larger and more variable effect estimates. Slope for C-reactive protein and RES 1.14 (-0.06, _2.35) and PERS 1.67 (-0.03, 3.37). Slope for Interleukin-6 and RES 0.35 (-0.28, 1.34) and PERS 0.76 (-0.36, 1.88). Explained largely by smaller contrast in PERS exposure	[25]
Canchec 2001 census	Nationwide in Canada	Time weighted average of actual work and home location	PM <sub>2.5</sub> at 1*1 km, MAPLE satellite -based method	Nearly identical associations with mortality. In the full cohort, HR were 1.12 (1.10, 1.15) and 1.11 (1.09, 1.13) for a 10 µg/m3 increment for RES and PERS respectively. In the cohort that reported regular commute to work HR were 1.17 (1.07, 1.28) and 1.15 (1.04, 1.27) for RES and PERS. Exposure at work was not associated with mortality, HR = 1.01 (0.93, 1.09)	E
Children 8 yr of PlAMA birth cohort (N=668)	Nationwide in the Netherlands	ABM, incorporating school, commut- ing to school and playing outdoors. Time fractions and distance to loca- tions from Dutch surveys	$\rm NO_{2^{\prime}}PM_{25^{\prime}}PM_{2.5}$ absorbance from LUR	Nearly identical associations with lung function, e.g., per IOR increase in NO <sub>2</sub> with RES -1,4% (-2.5, -0.2) and PERS with estimated school -1.6% (-2.8, -0.4) and -1.5 (-2.7, -0.3) actual school	[36]
Tracking panel survey of adults (N = 393)	The Netherlands	Activity-space based assessment inte- grating actual GPS readings for 7 days	PM <sub>2.5</sub> from LUR	No significant association with depres- sive symptoms for residential and dynamic exposure. For the 100-m buffer, regression slopes of the fully adjusted model were 0.225 (0.384) and 0.183 (0.472) for RES and PERS	[47]
Adults in San Diego (V = 552)	San Diego, USA	Tracking	$NO_{2^{\prime}}$ PM <sub>2.5</sub> using interpolation of routine monitoring data	No association with static exposure (1600 m buffer) and significant asso- ciations with some cardio-metabolic outcomes e.g., for NO2 and PM2.5 RR for metabolic syndrome were 0.92 (0.84, 1.02) and 1.07 (0.98, 1.18) for RES and 1.17 (1.07, 1.28) and 1.21 (1.12, 1.30) for dynamic, respectively	Letellier, 2022 [27]

activity integrated exposure only and timeresidential address at the đ estimates in enidemininal studies using Table 2 Studies comparing health effect mortality were found with both residential and timeweighted average exposure of residential and work address [7]. The actual work address was reported in the census questionnaire. Hazard ratios (HRs) were almost the same for the two exposure metrics, both in the full cohort and the cohort of people reporting that they commuted to work. Furthermore, no association with mortality was found when only the exposure at the work address was used. This observation agreed with a study in the Health Professionals cohort in the USA, in which no association with mortality was found for work address PM<sub>25</sub> [44]. The lack of association with air pollution at the work address in two studies could be related to a combination of the reduced time spent at work or less infiltration of ambient particles into the occupational environment. A study in Hong-Kong indeed reported lower infiltration ratios for commercial buildings than for residences and schools [58]. For  $PM_{2.5}$ , infiltration ratios were 0.82, 0.40 and 0.92 for residences, commercial buildings and schools, respectively. For NO<sub>2</sub> differences in infiltration were much smaller [58]. More work is needed to assess how general this finding is.

In a study in 700 Dutch primary school children aged 8 years, associations between long-term exposure to  $PM_{2.5}$ ,  $NO_2$  and  $PM_{2.5}$  absorbance with lung function were similar for residential and time-activity integrated exposure, based on agent-based modelling [36]. Health effect estimates were nearly identical when the estimated and the actual school location were used in the exposure assessment.

In a study of 500 adults in San Diego, USA,  $PM_{2.5}$  and  $NO_2$  were associated with some cardio-metabolic markers for dynamic exposures only [27]. For the residential exposure, no associations were found. However, exposure was characterized as the average exposure in a 1600 m buffer around the address, thus neglecting fine-spatial resolution variations in outdoor concentrations. As the actual exposure to outdoor generated air pollution largely occurs indoors, the focus on characterizing residential

derived from a detailed (15-min resolution) time-activity diary [25].

In a cohort study of 2,497 kindergarten and first-grade children of the Southern California Children's Health Study (USA), incident asthma was associated with both home and school exposure to traffic-related air pollution obtained from a line source dispersion model [33]. A time-weighted average exposure at home and school gave almost identical relative risks compared to the residential exposure.

In a tracking study of 392 Dutch adults (18–65 years), sampled from a large representative survey by Statistics Netherlands, depressive symptoms from a standard questionnaire were not significantly associated with  $PM_{2.5}$  exposure estimated at the residential address and with mobility integrated [47]. The authors did find a significant negative association with green space, both for residential and mobility-integrated exposure [47].

Two further studies combined school and residential exposure, but did not report a comparison with residential only exposure [17, 34]. Both studies did not find an association with asthma.

# Bias in epidemiological studies calculated from exposure validation studies

Table 3 shows the six studies that estimated the potential bias related to assessing exposure at the residential address only. Consistent with the high correlation and the small differences in effect estimates found in epidemiological studies, these six studies generally showed small to moderate bias in estimated exposure response functions, assuming classical error [11, 28, 40, 49] (Ragettli, 2015; Yu, 2020). The specific formula used in these studies was:

BIAS factor = 
$$(\sigma^2 + \rho)/(\sigma^2 + 2\rho + \omega^2)$$

where

 $\sigma^2$  = variance of the true (time – activity integrated) exposure

 $\rho = \text{ covariance of residential and time} - \text{ activity integrated exposure}$ 

 $\omega^2 =$  variance of the error (difference between residential and time-activity integrated)

outdoor air pollution in the immediate vicinity to the home, instead of a large buffer around it, in virtually all epidemiological studies seems more appropriate.

In a study in 140 adults living near a major interstate in the Greater Boston area (USA) or in a control neighborhood farther away from the interstate, associations of long-term exposure to UFP with the inflammation markers C-Reactive Protein and Interleukin-6 were weak with both home-based and mobility-integrated exposures The bias factor expresses the degree to which a health effect is underestimated, e.g., a bias factor of 0.7 implies that the theoretical true risk is underestimated by 30%.

The Setton [49] study evaluated potential bias by comparing residential address only  $NO_2$  exposure and an indirect personal exposure integrating work and commuting exposures using simulations in Vancouver based on a Canadian time-activity survey and actual individual-level time-activity data in a Southern

Population	Study area	Dynamic EA method	Pollutants	Bias factor	Reference
Vancouver (n = 386 census tracts) and US survey participants (N = 18,547)	Vancouver, Canada and Southern California, USA	Based on survey data, origin destination matrices (Vancouver) and actual addresses and trips (South-California). Work address and commuting included	NO <sub>2</sub> assessed with LUR (both areas) and Inverse distance weighing (IDW, Vancouver)	Mean bias factor in Vancouver 0.70 for LUR and 0.84 for IDW. Mean bias 0.93 for the full South-Cali- fornia population. Bias increases with time away from home and distance between work and home. Bias less than 10% when 0–4 h per day away from home and work>40 km from home: 25% when 4–10 h; 40% when>10 h/day; bias smaller for shorter distance	[49]
Participants Swiss time survey (N= 680 adults and children)	Cantons of Basel-Stadt and Baselland, Switzerland	Actual work addresses and geo- coded trips	NO <sub>2</sub> assessed with 100*100 m resolution dispersion model	Bias factor 0.87 when compared with home, work and commut- ing; bias 0.88 when compared with work address only	Ragettli, 2014 [45]
Mobile phone users of two companies (N=407,435)	Greater Boston area, USA	Anonymous mobile phone data used to infer home and work location	$PM_{2.5}$ from spatio-temporal satellite founded LUR model	Bias factor 0.91 comparing home and home-work combined estimate	[40]
Random sample of residents (N=100,789)	Los Angeles County, USA	Agent-based modeling on time survey data in Southern-California	$PM_{2.5}$ at 500*500 m resolution	Bias factor 0.87 for full population (0.78 for workers and 0.95 for non- workers)	[28]
Sample of workers from adminis- trative database (N= 130 million)	United States	Combination of home and work address	$PM_{2,5}$ from remote sensing at 1.1 * 1.1 km resolution	Bias factor 0.96 – 0.98, depend- ing on race, income, age and urbanicity	[11]
Adults (N = 310,989)	Shenzhen, China	Cell-phone location data	PM <sub>25</sub> , NO <sub>2</sub> , CO, NO2, SO2, at 3*3 km by CMAQ and inverse distance weighing	Bias factor for NO <sub>2</sub> and PM <sub>2.5</sub> dependent on mobility. Bias factor between 0.85 and 1.0 (NO <sub>2</sub> ) and 0.83 and 1.0 (NO <sub>2</sub> ) and 0.83 and 1.0 (PM <sub>2.5</sub> ) for 95% of the population. Lowest bias factor 0.67 and 0.70 for highest mobility group (0.5% of the population)	Yu, 2018 [62]

Table 3 Studies assessing potential bias by using exposure at the residential address only compared to time-activity integrated exposure

California time survey of 25,000 adults [49]. In the Vancouver analysis, 10,000 random draws of single locations from each of 382 census tracts were performed. In both settings, subjects who did not travel, or travelled within the same census tract, were not included in the analysis. This implies that the estimated bias overestimates the actual bias in a mixed population of commuters and non-commuters. Overall, the bias in Vancouver was moderate. However, the bias factor was larger when a more spatially varying surface was included in the analysis (0.70 for LUR compared to 0.84 for interpolated surface). The bias factor in Southern California was small (mean 0.93) but increased for those travelling longer distances from home to work and working more hours (Table 3). For those working 6-8 h and more than 40 km away from home, the bias factor was 0.70 and higher.

The study in Basel-Stadt and Baselland extracted information on commute routes, home, work and school locations from geo-coded 24-h time-activity diaries from the 2010 Swiss Mobility and Transport Microcensus [45]. This national telephone-based survey includes coordinates of origin and destination locations and geo-coded travel, including travel modes, duration and hour of the day. In the sample population, 76% of time was spent at home and 22% at work/school. The mean commute duration and distance of the study participants was 49 min and 14 km per day. The bias factor for assessing exposure at the residential address only was 0.87 when compared with the integrated home, work and commuting exposures. The largest contribution was from the work address exposure, as the bias factor was 0.88 when commuting was excluded.

A study in the greater Boston area, estimated work and home address location from mobile phone data for over 400,000 individuals [40]. A comparison of home-based and home-work time weighted exposures of  $PM_{2.5}$ , linked from a fine resolution spatio-temporal LUR model, showed nearly identical mean exposures with moderately lower contrast in home-work exposure. The bias factor calculated assuming classical error was 0.91.

A study in Los Angeles County used agent-based modelling combined with  $PM_{2.5}$  concentrations estimated at 500×500 m resolution for a group of about 100,000 individuals [28]. A modest bias factor of 0.87 was found [28]. The bias factor indicated more bias for commuters (0.78) than for non-working people (0.95).

A study in Shenzhen using cell-phone data combined with exposure modelled at 3\*3 km resolution, found that bias factor for NO<sub>2</sub> and PM<sub>2.5</sub> varied between 0.70 and 1 dependent on mobility across grids [62]. Bias factors below 0.8 were calculated for less than 5% of the population.

The six studies document small to moderate potential bias in epidemiological studies. We note that the impact of measurement error on effect estimates in reality is more complex and depends, for example, on the contribution of classical and Berkson error to the total error [48]. We thus view these calculations as an illustration of the potential bias. For the interpretation of these findings, the following observations are important.

First, the modelling and epidemiological studies that have integrated work or school address and commuting in the assessment of personal exposure have not considered all activity patterns, e.g. leisure time activities. These activities are more difficult to model. However, work and school are likely to be the key micro-environments beyond the home for a large part of the population (children and non-retired adults), based on time survey data. A study in Israel documented the largest discrepancy between residential and time-activity integrated exposure related to exposure during work time [51]. Furthermore, the studies based upon actual tracking and agent-based modelling including leisure time have found similar high correlations and effect estimates.

Second, in the current assessment, we have devoted little attention to indoor exposures. If the findings in the Hong Kong study of less infiltration of outdoor particles in the indoor environment of workplaces are more widely applicable [58], the residential environment would be even more important than the work environment for most (indoor) occupations.

Third, the potential bias may be outcome or age-group dependent, for example on whether we study childhood disease or adult morbidity or mortality. For the large body of literature on mortality effects of air pollution, the agreement for the elderly needs to be assessed. Residential exposure may be more closely related to personal exposure for this age group, though we lack the data to support this hypothesis.

### Implications, generalizability and future work

The main implication of the high correlation between residence-based and estimated personal exposure is that epidemiological studies on long-term outdoor air pollution exposure are unlikely to have substantial bias related to exposure assessment. In particular, studies are unlikely to miss true associations. Consistently, six studies showed only moderate bias in effect estimates related to assumed classical exposure measurement error (Table 3). More directly, in five epidemiological studies the differences in effect estimates were small between residential and time-activity integrated personal exposure [7, 25, 33, 36, 47].

Effect estimates may differ if there is a sizable difference in exposure contrast between residential and dynamic exposure assessments, even if the correlation between the two metrics is high. However, most studies showed only a moderately smaller contrast in exposure for time-activity integrated exposure compared to residential exposure.

Because of the observation that differences between residential and personal exposure may differ between subgroups of the population, it is possible that subgroup analysis in epidemiological studies may be more affected than the overall effect estimate. No studies have empirically evaluated this. The same applies to health impact assessment based upon epidemiological evidence.

High correlations between residential and dynamic exposure have been found in a diversity of populations, in children and adults and in different countries. However, as illustrated for a Californian and Canadian population, the potential bias depends on specific characteristics of the population studied, including time spent out of the home, distance between home and work and the spatial variability of outdoor air pollution [49]. In the large Canadian cohort study, a distinction was made between commuting and non-commuting subjects (Christidis, 2019). Bias may be larger in settings with longer times spent out of the home and larger distances between home and work than in the studies evaluated in this review. For example, in specific populations such as truck drivers or aircraft personnel, the bias may be more substantial. For children with a long commute in previously polluting diesel-powered school busses, bias may also be more substantial than has been reported for the Dutch studies of children visiting schools in their own neighborhood. Hence, in cohort studies it is worthwhile to collect information on work / school address, commuting times and mode of transport. Tracking studies in smaller populations also remain useful for validation. It remains difficult to imagine the application of dynamic exposure assessment in the mega-cohorts based on administrative data [13, 55].

Despite the consistency across different study designs (exposure, epidemiology), further work on the comparison of health effect estimates in epidemiological studies is useful, particularly in different populations than those studied to date. In the studies calculating bias based on exposure comparisons, more sophisticated analyses, accounting for both classical and Berkson error would be useful.

The observed usefulness of the residential exposure furthermore does not imply that we should not address air pollution exposures elsewhere. Several studies in children have reported health effects related to outdoor air pollution exposure at school, including studies in the Netherlands [6, 21], Barcelona [1, 43, 57] and California [33]. These studies specifically measured air pollution exposures at and in the schools. In one study in children [43], inflammation markers and blood pressure were associated both with home and school-based exposure in the week prior to the health evaluation.

In population exposure assessment, it is likely useful to account for more than just the residential address. Studies showed that differences in exposure between different socio-economic groups were more pronounced for personal compared to residential exposure [3]. Hence in evaluating environmental injustice there is benefit in a broader exposure assessment approach. The level of exposure also generally differs between residential and personal exposure. It is not obvious that this in itself is of great interest as the judgement of health risks can only be based on residential outdoor levels, as this is the method used in the epidemiological studies forming the evidence base to assess health risks. When comparing different populations, time-integrated exposure may show a different pattern if people predominantly work in differently exposed locations e.g. people living in suburbs / rural towns working either in a major city center versus more rural places. Finally, in studies on the impact of interventions on commuting exposures will be less reflected at the residential address.

## Conclusions

We found that the correlation between residential address-based and dynamic time-activity integrated long-term air pollution exposure was generally high to very high in diverse populations. Consistent with the high correlation, epidemiological studies found generally similar health effects using residential and dynamic exposure. Six additional studies showed only small to moderate bias in estimated exposure response functions using residence-based exposures, assuming classical error. Overall, the bias in epidemiological studies related to assessing long-term exposure at the residential address only is likely small in populations comparable to those evaluated in the reviewed studies.

Exposure studies did show generally modest differences in exposure level and exposure contrast, potentially affecting the magnitude of estimated health effects. The difference in exposure, however, is not equally distributed across the population including between different socio-economic groups, urban and rural populations and different neighborhoods. Further improvements in exposure assessment for epidemiological and health impact assessment studies, especially for large populations, remain highly useful. A limitation of the studies is that few considered indoor – outdoor relationships. Incorporation of infiltration of outdoor air pollutants may further improve exposure assessment.

### Abbreviations

GPS Geographical positioning syste
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- HR Hazard ratio
- NO<sub>2</sub> Nitrogen dioxide
- $PM_{2.5}$  Fine particles, particles smaller than 2.5  $\mu m$
- R Correlation coefficient
- UFP Ultrafine particles, particles smaller than 100 nm

### Authors' contributions

All authors contributed to the conceptualization of the study. GH and KH performed the search for studies and prepared the tables. All authors critically read and approved the final manuscript.

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#### Availability of data and materials

No datasets were generated or analysed during the current study.

### Declarations

Ethics approval and consent to participate

Not applicable.

### Competing interests

The authors declare no competing interests.

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